

SHIP/KEF-Richtlinien zur Durchführung qualitäts- gesicherter Auswertungen

1. Für Auswertungsprojekte ist im Regelfall ein bewilligter Datennutzungsantrag erforderlich.
2. Vor Beginn der Analysen ist ein Analyseplan zu erstellen. Dieser ist zumindest zwischen Erst-, Zweit- und Letztautor abzustimmen.
3. Variablendefinitionen und statistische Analysen sind vollständig in Programmen (z.B. Stata-Do-Files und Log-Files) zu dokumentieren. Diese Programme sind vor Ersteinreichung des Manuskripts an den Leiter des Publikationskomitees zu schicken. (siehe Appendix mit Beispielanalysen)
4. Die Log-Files, die alle Variablendefinitionen und statistischen Analysen beinhalten, sind durch eine zweite Person zu kontrollieren. Hierzu erfolgt insbesondere ein Abgleich mit den Tabellen des Manuskripts.
5. Die Analysen sind sachgerecht durchzuführen, dies umfasst:
 - die Auswertungen mit adäquaten statistischen Verfahren und
 - die Überprüfung der Annahmen der verwendeten Verfahren [1,2,3]. Bei erheblicher Verletzung der Annahmen sind die Analysen entsprechend anzupassen. Die Überprüfung der Annahmen ist ebenfalls im Do- und Log-File zu dokumentieren.

Appendix

Allgemeine Hinweise

In der Regel stehen die Ergebnisse von Regressionsmodellen in einem Auswertungsprojekt im Vordergrund. Bei der Schätzung von Regressionsmodellen ist sicherzustellen, dass die zugrundeliegenden Annahmen erfüllt sind [1,2,3]. Substantielle Verletzungen dieser Annahmen können zu Verzerrungen der Schätzer (Koeffizienten, P-Werte, Konfidenzintervalle, etc.) führen.

Wichtige Annahmen im Rahmen klassischer Regressionsmodellen umfassen:

- a. Normalverteilung der Residuen
- b. Linearität
- c. Homoskedasizität
- d. Einflussreiche Beobachtungen
- e. Unabhängigkeit der Residuen
- f. Missing completely at random (MCAR) Annahme bei fehlenden Werten und Ausfällen

Hinweise zur Kovariablen-Auswahl

Die Berücksichtigung und Auswahl von Kovariablen in einem Regressionsmodell kann unterschiedliche Ziele verfolgen. Grundsätzlich ist zwischen Konfounder- und Prädiktionsmodellen zu differenzieren [1,2,3]. Im Konfoundermodell steht der Zusammenhang zwischen Exposure und Outcome im Vordergrund. Die Auswahl zu adjustierender Konfounder erfolgt theoriegeleitet (z.B. mittels kausaler Grafen) oder anhand empirischer Kriterien (z.B. 15%-Change-in-Coefficient) [4]. Bei der Entwicklung von Prädiktionsmodellen [3] kommen die theoriegeleitete und empirisch-automatisierte Selektion (wie Stepwise-Verfahren) zur Anwendung [3].

Quellen:

1. Vittinghoff et al 2005. Regression Methods in Biostatistics.
2. Harrell 2001. Regression Modeling Strategies.
3. Steyerberg 2010. Clinical Prediction Models.
4. Rothman, Grenland, Lash 2008. Modern Epidemiology.

Beispielsyntax und Log-File für die Durchführung und Überprüfung einer Regressionsanalyse

Das folgende Beispiel dient zur Illustration von Do- und Log-Dateien. Analysiert werden Korrelate des Geburtsgewichts mittels linearer Regression.

1. Log-File zur Variablendefinition

```
*****
. ** Sebastian Baumeister
. ** Titel of project: Maternal correlates of birth weight
. ** Pfad: E:\Arbeit\Greifswald\ship_statistik\Richtlinie zur Durchführung von Auswertungen\
*****
```

```
glo bw "E:\Arbeit\Greifswald\ship_statistik\Richtlinie zur Durchführung von
Auswertungen\data\"
```

```

. *Load original dataset and save under new name
. *****
. use http://www.stata-press.com/data/r10/lbw3.dta, clear
(Hosmer & Lemeshow data)

save "$bw\bw_smoking.dta", replace
file E:\Arbeit\Greifswald\ship_statistik\Richtlinie zur Durchführung von
Auswertungen\data\bw_smoking.dta saved

codebook, c
Variable  Obs Unique      Mean  Min  Max  Label
-----
id        189    189  121.0794   4  226  identification code
low       189     2   .3121693   0   1  birth weight<2500g
age       189    24   23.2381  14   45  age of mother
lwt       189    76  129.8201  80  250  weight at last menstrual period
race      189     3   1.846561   1   3  race
smoke     189     2   .3915344   0   1  smoked during pregnancy
ptl       189     4   .1957672   0   3  premature labor history (count)
ht        189     2   .0634921   0   1  has history of hypertension
ui        189     2   .1481481   0   1  presence, uterine irritability
ftv       189     6   .7936508   0   6  number of visits to physician during 1st trimester
bwt       189   133  2944.286  709 4990  birth weight (grams)
white     189     2   .5079365   0   1  race==white
black     189     2   .1375661   0   1  race==black
other     189     2   .3544974   0   1  race==other
ptd       189     2   .1587302   0   1
lwd       189     2   .2222222   0   1
* Define new outcome variables
. *****
Low birth weight
recode bwt (min/2500=1) (2500.001/max=0), gen(lowbwt)
(189 differences between bwt and lowbwt)
label variable lowbwt "birth weight"
label define lowbwt 1 "1,birth weight<2500g" 0 "0,birth weight>=2500"
label value lowbwt lowbwt

. *Covariables
. *****
. * Age groups
. recode age (min/19=1) (20/23=2) (24/26=3) (27/max=4) (.=.), gen(age4)
(189 differences between age and age4)

. save, replace
file E:\Arbeit\Greifswald\ship_statistik\Richtlinie zur Durchführung von
Auswertungen\data\bw_smoking.dta saved

. log close

```

Der 2. Log-File unten veranschaulicht das Vorgehen bei der Durchführung und Überprüfung der Annahmen der linearen („kleinste Quadrate“) Regression.

2. Log-File zur Auswertung

```

.
*****
*****
. ** Sebastian Baumeister
. ** Titel of project: Maternal correlates of birth weight
. ** Pfad: E:\Arbeit\Greifswald\ship_statistik\Richtlinie zur Durchführung von Auswertungen\
*****
*****
glo bw "E:\Arbeit\Greifswald\ship_statistik\Richtlinie zur Durchführung von
Auswertungen\data\"

. use "$bw\bw_smoking.dta", clear
(Hosmer & Lemeshow data)

des, sh

Contains data from E:\Arbeit\Greifswald\ship_statistik\Richtlinie zur Durchführung von
Auswertungen\data\bw_smoking.dta
   obs:          189              Hosmer & Lemeshow data

```

```
vars:          18                      10 Jan 2012 14:15
size:         5,292
Sorted by:
```

```
. * Analytical sample
. *****
. glo out bwt
```

```
. glo cov age lwt race smoke ptl ht ui ftv
```

```
. tabmiss $out $cov
```

Variable	Obs	Missings	Freq.Missings	NonMiss	Freq.NonMiss
bwt	189	0	0	189	100
age	189	0	0	189	100
lwt	189	0	0	189	100
race	189	0	0	189	100
smoke	189	0	0	189	100
ptl	189	0	0	189	100
ht	189	0	0	189	100
ui	189	0	0	189	100
ftv	189	0	0	189	100

```
. egen nm=rowmiss($out $cov)
```

```
. fre nm
```

```
nm
```

	Freq.	Percent	Valid	Cum.
Valid 0	189	100.00	100.00	100.00

```
. recode nm (0=0) (1/max=1),gen(miss)
(0 differences between nm and miss)
```

```
. * No missing values on any variable included in the (regression analyses)
```

```
. * Table 1: Characteritics of the mothers
```

```
. *****
```

```
. tabstat age lwt ptl ftv if miss==0, s(n median p25 p75 mean sd) c(s)
```

variable	N	p50	p25	p75	mean	sd
age	189	23	19	26	23.2381	5.298678
lwt	189	121	110	140	129.8201	30.57515
ptl	189	0	0	0	.1957672	.4933419
ftv	189	0	0	1	.7936508	1.059286

```
. fre race smoke ht ui if miss==0
```

```
race -- race
```

	Freq.	Percent	Valid	Cum.
Valid 1 white	96	50.79	50.79	50.79
2 black	26	13.76	13.76	64.55
3 other	67	35.45	35.45	100.00
Total	189	100.00	100.00	

```
smoke -- smoked during pregnancy
```

	Freq.	Percent	Valid	Cum.
Valid 0	115	60.85	60.85	60.85
1	74	39.15	39.15	100.00
Total	189	100.00	100.00	

```
ht -- has history of hypertension
```

	Freq.	Percent	Valid	Cum.
--	-------	---------	-------	------

Valid	0		177	93.65	93.65	93.65
	1		12	6.35	6.35	100.00
Total			189	100.00	100.00	

ui -- presence, uterine irritability

		Freq.	Percent	Valid	Cum.
Valid	0		161	85.19	85.19
	1		28	14.81	100.00
Total			189	100.00	100.00

. * Table 2: Correlates of birth weight

. *****
. reg bwt age i.race i.smoke i.ht lwt ptl ftv if miss==0

Source	SS	df	MS	Number of obs =	189
Model	18427359.2	8	2303419.9	F(8, 180) =	5.09
Residual	81487939.3	180	452710.774	Prob > F =	0.0000
				R-squared =	0.1844
				Adj R-squared =	0.1482
Total	99915298.6	188	531464.354	Root MSE =	672.84

bwt	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
age	-1.400947	9.932138	-0.14	0.888	-20.99935	18.19745
race						
2	-488.0906	155.1772	-3.15	0.002	-794.291	-181.8901
3	-372.048	118.6461	-3.14	0.002	-606.1641	-137.932
1.smoke	-362.5319	110.121	-3.29	0.001	-579.826	-145.2379
1.ht	-524.4655	208.4684	-2.52	0.013	-935.8216	-113.1093
lwt	4.801987	1.791585	2.68	0.008	1.266777	8.337197
ptl	-127.4346	103.163	-1.24	0.218	-330.999	76.12981
ftv	-9.086507	48.06139	-0.19	0.850	-103.9227	85.7497
_cons	2759.884	320.2368	8.62	0.000	2127.982	3391.785

. * Test of linear regression (OLS) assumptions

. *****
. * Normality of residuals: looks fine

. *****
. predict r, res

. kdensity r, norm

. qnorm r

. pnorm r

. * Homoscedasticity: looks fine

. *****
. estat hettest age i.race i.smoke i.ht lwt ptl ftv, iid

Breusch-Pagan / Cook-Weisberg test for heteroskedasticity

Ho: Constant variance

Variables: age i.race i.smoke i.ht lwt ptl ftv

chi2(8) = 7.45

Prob > chi2 = 0.4887

. rvfplot, yline(0)

. * Linearity

. *****

. * Age - looks nonlinear, modeled using restricted cubic splines

. twoway (scatter bwt age) (lfit bwt age) (lowess bwt age) (fpfit bwt age)

. acprplot age, lowess lsopts(bwidth(1))

```
. centile age if miss==0, c(5 35 65 95)
```

Variable	Obs	Percentile	Centile	-- Binom. Interp. -- [95% Conf. Interval]	
age	189	5	16	15	17
		35	20	20	21
		65	25	24	26
		95	32.5	31	35

```
. mkspline2 ag_age, cubic knots(16 20 25 32.5) dis
```

	knot1	knot2	knot3	knot4
age	16	20	25	32.5

```
. * Weight at last menstrual period - looks linear
. twoway (scatter bwt lwt) (lfit bwt lwt) (lowess bwt lwt) /*(fpfit bwt lwt)*/
```

```
. * premature labor history
. twoway (scatter bwt ptl) (lfit bwt ptl) (lowess bwt ptl) /*(fpfit bwt lwt)*/
```

```
. fre ptl
```

```
ptl -- premature labor history (count)
```

		Freq.	Percent	Valid	Cum.
Valid	0	159	84.13	84.13	84.13
	1	24	12.70	12.70	96.83
	2	5	2.65	2.65	99.47
	3	1	0.53	0.53	100.00
Total		189	100.00	100.00	

```
. *only few subjects with premature labor history >1 therefore recode with new variables values 0 and 1+
```

```
. recode ptl (0=0 "0,zero") (1/3=1 "1,1+"), gen(plt_2c)
(6 differences between ptl and plt_2c)
```

```
. * Collinearity - not an issue
```

```
. *****
. vif
```

Variable	VIF	1/VIF
age	1.15	0.869446
race		
2	1.19	0.838428
3	1.34	0.743603
1.smoke	1.21	0.829112
1.ht	1.08	0.926932
lwt	1.25	0.802511
ptl	1.08	0.929649
ftv	1.08	0.929060
Mean VIF	1.17	

```
. collin bwt age race smoke ht lwt plt_2c ftv if miss==0
(obs=189)
```

Collinearity Diagnostics

Variable	VIF	SQRT VIF	Tolerance	R- Squared
bwt	1.22	1.10	0.8211	0.1789
age	1.13	1.06	0.8848	0.1152
race	1.30	1.14	0.7684	0.2316
smoke	1.28	1.13	0.7816	0.2184
ht	1.12	1.06	0.8955	0.1045
lwt	1.20	1.09	0.8364	0.1636
plt_2c	1.11	1.05	0.9010	0.0990
ftv	1.07	1.04	0.9304	0.0696
Mean VIF	1.18			

	Eigenval	Cond Index
1	5.7714	1.0000
2	0.9465	2.4693
3	0.8763	2.5664
4	0.5870	3.1356
5	0.5637	3.1997
6	0.1628	5.9545
7	0.0438	11.4792
8	0.0373	12.4466
9	0.0112	22.6979

Condition Number 22.6979
Eigenvalues & Cond Index computed from scaled raw sscp (w/ intercept)
Det(correlation matrix) 0.5271

. * Omitted variables / exogeneity -> not suspected
. *****
. linktest

Source	SS	df	MS	Number of obs =	189
Model	18566610.8	2	9283305.38	F(2, 186) =	21.23
Residual	81348687.8	186	437358.537	Prob > F =	0.0000
				R-squared =	0.1858
				Adj R-squared =	0.1771
				Root MSE =	661.33
Total	99915298.6	188	531464.354		

bwt	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
_hat	2.244969	2.211734	1.02	0.311	-2.118339	6.608278
_hatsq	-.0002112	.0003743	-0.56	0.573	-.0009497	.0005272
_cons	-1813.978	3246.971	-0.56	0.577	-8219.603	4591.646

. estat ovtest

Ramsey RESET test using powers of the fitted values of bwt
Ho: model has no omitted variables
F(3, 177) = 1.49
Prob > F = 0.2184

. * Influential observations
. *****
. *Cookd distance > 4/N
. reg bwt ag_* i.race i.smoke i.ht lwt ptl ftv if miss==0

Source	SS	df	MS	Number of obs =	189
Model	20672624.4	10	2067262.44	F(10, 178) =	4.64
Residual	79242674.1	178	445183.563	Prob > F =	0.0000
				R-squared =	0.2069
				Adj R-squared =	0.1623
				Root MSE =	667.22
Total	99915298.6	188	531464.354		

bwt	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
ag_1	16.37202	49.33889	0.33	0.740	-80.9924	113.7364
ag_2	-232.9313	218.3098	-1.07	0.287	-663.7397	197.877
ag_3	612.1989	476.4075	1.29	0.200	-327.9347	1552.332
race						
2	-495.2098	154.1368	-3.21	0.002	-799.3804	-191.0392
3	-353.5734	117.9436	-3.00	0.003	-586.321	-120.8258
1.smoke	-369.0548	109.3287	-3.38	0.001	-584.802	-153.3076
1.ht	-511.5395	207.5715	-2.46	0.015	-921.1572	-101.9219
lwt	4.648354	1.7802	2.61	0.010	1.135341	8.161367
ptl	-104.6919	102.8709	-1.02	0.310	-307.6954	98.31169
ftv	-10.29435	47.67845	-0.22	0.829	-104.3821	83.79341
_cons	2535.593	906.5436	2.80	0.006	746.6368	4324.548

```
. predict d, cooks
```

```
. li bwt age race smoke ht lwt plt_2c ftv if d>4/189
```

	bwt	age	race	smoke	ht	lwt	plt_2c	ftv
23.	2836	36	white	0	0	202	0,zero	1
65.	3260	26	other	1	0	133	1,1+	0
94.	3637	25	white	1	0	95	1,1+	0
102.	3756	19	white	1	1	184	0,zero	0
106.	3790	25	black	0	1	241	0,zero	0
110.	3860	16	black	0	0	170	0,zero	4
130.	4990	45	white	0	0	123	0,zero	1
131.	709	28	other	1	0	120	1,1+	0
132.	1021	29	white	0	0	130	0,zero	2
133.	1135	34	black	1	1	187	0,zero	0
134.	1330	25	other	0	1	105	1,1+	0
147.	1928	21	black	0	0	200	0,zero	2

```
. *dfits > 2*sqrt(k/N)
```

```
. predict dfits, dfits
```

```
. quiet scalar thresh=2*sqrt((e(df_m)+1)/e(N))
```

```
. di "dfits threshold=" %6.3f thresh
```

```
dfits threshold= 0.482
```

```
. li dfits bwt age race smoke ht lwt plt_2c ftv if abs(dfits)>2*thresh & e(sample)
```

	dfits	bwt	age	race	smoke	ht	lwt	plt_2c	ftv
94.	1.032447	3637	25	white	1	0	95	1,1+	0
130.	1.464016	4990	45	white	0	0	123	0,zero	1
133.	-.9895521	1135	34	black	1	1	187	0,zero	0

```
. reg bwt ag_* i.race i.smoke i.ht lwt plt ftv if miss==0 & d<4/189
```

Source	SS	df	MS	Number of obs =	177
Model	17613347.8	10	1761334.78	F(10, 166) =	4.90
Residual	59651728.5	166	359347.762	Prob > F =	0.0000
				R-squared =	0.2280
				Adj R-squared =	0.1815
Total	77265076.3	176	439006.115	Root MSE =	599.46

bwt	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
ag_1	22.39731	46.39013	0.48	0.630	-69.19341 113.988
ag_2	-215.2098	216.1807	-1.00	0.321	-642.0279 211.6083
ag_3	569.4761	486.5547	1.17	0.244	-391.157 1530.109
race					
2	-441.4639	146.7486	-3.01	0.003	-731.1982 -151.7296
3	-320.2998	109.1899	-2.93	0.004	-535.8797 -104.72
1.smoke	-366.6639	101.9441	-3.60	0.000	-567.938 -165.3898
1.ht	-512.8817	221.7429	-2.31	0.022	-950.6815 -75.08184
lwt	4.619319	1.707276	2.71	0.008	1.248546 7.990092
plt	-225.7582	106.8228	-2.11	0.036	-436.6646 -14.85177
ftv	-7.792252	45.2459	-0.17	0.863	-97.12384 81.53934
_cons	2396.499	843.6123	2.84	0.005	730.9063 4062.091

```
. reg bwt ag_* i.race i.smoke i.ht lwt plt ftv if miss==0 & abs(dfits)<2*thresh
```

Source	SS	df	MS	Number of obs =	186
Model	17586801.3	10	1758680.13	F(10, 175) =	4.14
Residual	74385541	175	425060.234	Prob > F =	0.0000
				R-squared =	0.1912


```
-----+-----
Total | 91972342.3  185  497147.796
Adj R-squared = 0.1450
Root MSE      = 651.97
```

bwt	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
ag_1	5.318512	49.27923	0.11	0.914	-91.93959	102.5766
ag_2	-149.2398	226.7889	-0.66	0.511	-596.8333	298.3537
ag_3	405.2366	507.4508	0.80	0.426	-596.2747	1406.748
race						
2	-424.4026	153.8329	-2.76	0.006	-728.0091	-120.7961
3	-310.511	116.1151	-2.67	0.008	-539.6773	-81.34473
1.smoke	-324.2243	108.012	-3.00	0.003	-537.3981	-111.0505
1.ht	-403.0304	210.9367	-1.91	0.058	-819.3377	13.27692
lwt	5.147605	1.756653	2.93	0.004	1.680652	8.614557
ptl	-208.9705	109.7989	-1.90	0.059	-425.6709	7.729915
ftv	-9.28851	46.96304	-0.20	0.843	-101.9754	83.39833
_cons	2615.481	896.9764	2.92	0.004	845.1968	4385.764

```
. *Some influential observation that distort estimates
. *Robust and quantile regression
. rreg bwt ag_* i.race i.smoke i.ht lwt ptl ftv
```

```
Robust regression
Number of obs = 189
F( 10, 178) = 4.44
Prob > F = 0.0000
```

bwt	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
ag_1	14.28029	51.93272	0.27	0.784	-88.20275	116.7633
ag_2	-206.659	229.7867	-0.90	0.370	-660.1157	246.7977
ag_3	555.4032	501.4531	1.11	0.270	-434.1548	1544.961
race						
2	-485.2457	162.24	-2.99	0.003	-805.4071	-165.0843
3	-350.0935	124.1441	-2.82	0.005	-595.0771	-105.1099
1.smoke	-379.7665	115.0763	-3.30	0.001	-606.8559	-152.6771
1.ht	-503.1772	218.4839	-2.30	0.022	-934.3291	-72.0252
lwt	4.742754	1.873788	2.53	0.012	1.045056	8.440453
ptl	-133.221	108.2791	-1.23	0.220	-346.8968	80.45477
ftv	-8.67221	50.185	-0.17	0.863	-107.7063	90.3619
_cons	2564.002	954.2022	2.69	0.008	680.9976	4447.006

```
. xi: greg bwt ag_* i.race i.smoke i.ht lwt ptl ftv
i.race      _Irace_1-3      (naturally coded; _Irace_1 omitted)
i.smoke     _Ismoke_0-1    (naturally coded; _Ismoke_0 omitted)
i.ht        _Iht_0-1      (naturally coded; _Iht_0 omitted)
Iteration 1: WLS sum of weighted deviations = 98922.856
```

```
Median regression
Raw sum of deviations 111569 (about 2977)
Min sum of deviations 98122.23
Number of obs = 189
Pseudo R2 = 0.1205
```

bwt	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
ag_1	45.50179	92.97873	0.49	0.625	-137.9807	228.9842
ag_2	-272.5861	411.4234	-0.66	0.508	-1084.481	539.309
ag_3	663.8722	898.2262	0.74	0.461	-1108.67	2436.415
_Irace_2	-412.5939	279.1732	-1.48	0.141	-963.5091	138.3212
_Irace_3	-411.2986	219.5242	-1.87	0.063	-844.5035	21.90631
_Ismoke_1	-492.6311	203.2619	-2.42	0.016	-893.7442	-91.51801
_Iht_1	-459.1765	361.2127	-1.27	0.205	-1171.987	253.6339
lwt	5.314155	3.313364	1.60	0.111	-1.224375	11.85268
ptl	-172.1293	191.6626	-0.90	0.370	-550.3526	206.0939
ftv	-18.20978	88.36572	-0.21	0.837	-192.589	156.1694
_cons	1977.775	1705.814	1.16	0.248	-1388.446	5343.996

```

.
.
.
. *Final robust regression model
. xi: rreg bwt ag_* i.race i.smoke i.ht lwt i.plt_2c ftv if miss==0
i.race      _Irace_1-3      (naturally coded; _Irace_1 omitted)
i.smoke     _Ismoke_0-1    (naturally coded; _Ismoke_0 omitted)
i.ht        _Iht_0-1      (naturally coded; _Iht_0 omitted)
i.plt_2c    _Iplt_2c_0-1  (naturally coded; _Iplt_2c_0 omitted)

Robust regression                               Number of obs =      189
                                                F( 10,   178) =     4.79
                                                Prob > F       =     0.0000

```

	bwt	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
	ag_1	14.10818	51.27275	0.28	0.784	-87.07248	115.2888
	ag_2	-199.0313	227.0517	-0.88	0.382	-647.0908	249.0282
	ag_3	539.2802	495.4929	1.09	0.278	-438.516	1517.076
	_Irace_2	-469.5862	160.6347	-2.92	0.004	-786.5796	-152.5928
	_Irace_3	-335.1016	123.1111	-2.72	0.007	-578.0467	-92.15643
	_Ismoke_1	-358.5469	114.0825	-3.14	0.002	-583.6751	-133.4187
	_Iht_1	-487.7192	215.899	-2.26	0.025	-913.7703	-61.66819
	lwt	4.594307	1.848273	2.49	0.014	.9469598	8.241655
	_Iplt_2c_1	-285.7212	144.2481	-1.98	0.049	-570.3777	-1.06463
	ftv	-7.178501	49.54356	-0.14	0.885	-104.9468	90.58981
	_cons	2579.96	942.5491	2.74	0.007	719.9511	4439.968

```

. testparm ag_*

( 1) ag_1 = 0
( 2) ag_2 = 0
( 3) ag_3 = 0

      F( 3,   178) =     1.52
      Prob > F =     0.2110

. log close

```